



Exercice 1

$$1. f_{\theta}(x) = \frac{1}{\sqrt{2\pi\theta}} e^{-\frac{x^2}{2\theta}}$$

$$\text{V.S.: } L(\theta) = \prod_{i=1}^n f_{\theta}(x_i) = (2\pi\theta)^{-\frac{n}{2}} \exp\left(-\frac{1}{2\theta} \sum_{i=1}^n x_i^2\right)$$

$$\log \text{ V.S.: } l(\theta) = -\frac{n}{2} \ln(2\pi\theta) - \frac{1}{2\theta} \sum_{i=1}^n x_i^2$$

$$\begin{aligned} \text{dérivée: } l'(\theta) &= -\frac{n}{2\theta} + \frac{1}{2\theta^2} \sum_{i=1}^n x_i^2 \\ &= \frac{n}{2\theta^2} \left(\frac{1}{n} \sum_{i=1}^n x_i^2 - \theta \right) \end{aligned}$$

On résout (par rapport à θ):

$$l'(\theta) = 0$$

$$\Leftrightarrow \theta = \frac{1}{n} \sum_{i=1}^n x_i^2$$

$$\text{On pose } T = \frac{1}{n} \sum_{i=1}^n x_i^2$$

Vérifions qu'il s'agit d'un max global:

$$l'(\theta) = \frac{n}{2\theta^2} - \frac{1}{\theta^3} \sum_{i=1}^n x_i^2$$

$$l''(\theta) = \frac{n}{2\theta^3} - \frac{1}{\theta^4} \sum_{i=1}^n x_i^2 = \frac{n}{2\theta^3} - \frac{1}{\theta^3} = -\frac{n}{2\theta^3} < 0$$

Donc T est bien l'EMV de θ .

$$2. \text{ LFN: } \frac{1}{n} \sum x_i^2 \xrightarrow[n \rightarrow \infty]{p} E[X_i^2] = \theta$$

so

$$3. l(\theta)^2 = \left(\frac{n}{2\theta^2} \left(\frac{1}{n} \sum_{i=1}^n x_i^2 - \theta \right) \right)^2 = \frac{n^2}{4\theta^4} \left(\frac{1}{n^2} \left(\sum_{i=1}^n x_i^2 \right)^2 + \theta^2 - \frac{2}{n} \sum_{i=1}^n x_i^2 \right)$$

$$E[l'(\theta)^2] =$$

$$\begin{aligned} J_n(\theta) = \text{Var}(S(\theta)) &= \text{Var}(l'(\theta)) \\ &= \text{Var}\left(\frac{n}{2\theta^2} (T - \theta)\right) \end{aligned}$$

$$= \left(\frac{n}{2\theta^2}\right)^2 \text{Var}(T)$$

$$\begin{aligned} &= \left(\frac{n}{2\theta^2}\right)^2 \left(\frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i^2) \right) = \frac{n^2}{4\theta^4} \cdot \left(\frac{1}{n^2} \sum E[X_i^4] - E[X_i^2]^2 \right) \\ &= \frac{1}{4\theta^2} \cdot \left(\frac{1}{n^2} \sum 3\theta^2 - \theta^2 \right) = \frac{n^2}{4\theta^2} \cdot \frac{2\theta^2}{n} = \frac{n}{2\theta^2} \end{aligned}$$

la borne de Cramer-Rao pour tout estimateur sans-biais de θ si:

$$B C R(\theta) = \frac{1}{I_n(\theta)} = \frac{2\theta^2}{n}$$

Or T est sans-biais et $\text{Var}(T) = B C R(\theta)$. Donc T est un estimateur efficace de θ .

Exercice 2

$$1. \ell(\theta) = \prod_{i=1}^n f_{\theta}(x_i)$$

$$\begin{aligned} \log \ell(\theta) &= \sum_{i=1}^n \log f_{\theta}(x_i) \\ &= \sum_{i=1}^n \log\left(\frac{1}{\theta}\right) + \left(\frac{1}{\theta} - 1\right) \log x_i \\ &= n \log\left(\frac{1}{\theta}\right) + \underbrace{\left(\frac{1}{\theta} - 1\right)}_{\forall \theta > 0} \sum_{i=1}^n \log x_i \end{aligned}$$

$$\frac{\partial}{\partial \theta} \log \ell(\theta) = -\frac{n}{\theta} - \frac{1}{\theta^2} \sum_{i=1}^n \log(x_i) \quad \frac{\partial^2}{\partial \theta^2} \log \ell(\theta) = \frac{2n}{\theta^2} + \frac{2}{\theta^3} \sum \log(x_i)$$

$$\Leftrightarrow \frac{n}{\theta} = -\frac{1}{\theta^2} \sum_{i=1}^n \log(x_i)$$

$$\Leftrightarrow -\frac{n}{\sum \log(x_i)} = \frac{1}{\theta}$$

$$\Leftrightarrow \theta = -\frac{1}{n} \underbrace{\sum \log(x_i)}_{\leq 0} \geq 0$$

On pose $\hat{\theta}$, vérifications max glob.

$$\ell''(\theta) = \frac{n}{\theta^2} + \frac{2}{\theta^3} \sum_{i=1}^n \ln(x_i) = \frac{1}{\theta^3} \left(n\theta + 2 \sum_{i=1}^n \ln(x_i) \right)$$

$$\ell''(\hat{\theta}) = \frac{1}{\hat{\theta}^3} \left(n\hat{\theta}^2 + 2 \sum_{i=1}^n \ln(x_i) \right) = \underbrace{\frac{1}{\hat{\theta}^3}}_{\geq 0} \underbrace{\left(n\hat{\theta}^2 + 2 \sum_{i=1}^n \ln(x_i) \right)}_{\leq 0} \leq 0$$

$\hat{\theta}$ est bien l'EMV de θ .

$$2. Y_i = -\frac{1}{\theta} \log X_i$$

$$P\left(-\frac{1}{\theta} \log X_i < y\right)$$

$$P(\log X_i > -\theta y)$$

$$= P(X_i > e^{-\theta y})$$

$$= 1 - P(X_i \leq e^{-\theta y}) = 1 - \int_0^{e^{-\theta y}} \frac{1}{\theta} z^{\frac{1}{\theta}-1} dz$$

$$= 1 - \left[z^{\frac{1}{\theta}} \right]_0^{e^{-\theta y}} = 1 - e^{-\theta y \frac{1}{\theta}} = 1 - e^{-y} = F_{Y_i}(y)$$

donc $Y_i \sim \mathcal{E}(\theta)$

Loi de Z. Les X_i sont i.i.d. Les Y_i sont aussi:
donc $Z = \sum Y_i \sim \Gamma(n, \theta)$

$$\text{De plus } Z = -\frac{1}{\theta} \sum \ln(X_i) = \frac{n}{\theta} \hat{\theta} \Leftrightarrow \hat{\theta} = \frac{\theta}{n} Z$$

$$\text{Biais: } E[\hat{\theta}] = \frac{\theta}{n} E[Z] = \theta$$

donc $\hat{\theta}$ est sans biais.

$$3. \text{Var}(\hat{\theta}) = \left(\frac{\theta}{n}\right)^2 \text{Var}(Z) = \frac{\theta^2}{n}$$

$$4. \ell'(\theta) = -\frac{n}{\theta} - \frac{1}{\theta^2} \sum_{i=1}^n \ln(X_i) = \frac{n}{\theta^2} \left(-\theta - \frac{1}{n} \sum_{i=1}^n \ln X_i\right) \\ = \frac{n}{\theta^2} (\hat{\theta} - \theta)$$

5.

$$J_n(\theta) = \text{Var}(\ell'(\theta)) = \left(\frac{n}{\theta^2}\right)^2 \text{Var}(\hat{\theta}) = \frac{n^2}{\theta^4} \frac{\theta^2}{n} = \frac{n}{\theta^2}$$

$$\text{BCR} = \frac{1}{J_n(\theta)} = \frac{\theta^2}{n} = \text{Var}(\hat{\theta}) \quad \hat{\theta} \text{ est bien efficace.}$$

Exercice 3

$$\underline{\text{I}} \quad E[e^{tx}] = \int_{-\infty}^{+\infty} e^{tx} f_{\theta}(x) dx \\ = \int_0^{+\infty} e^{tx} \theta^{-2} x e^{-\frac{x}{\theta}} dx \\ = \theta^{-2} \int_0^{+\infty} x \exp(-x(\frac{1}{\theta} - t)) dx$$

On suppose $t < \frac{1}{\theta}$ sinon, l'intégrale diverge

$$\psi(t) = \theta^{-2} \left[\frac{-1}{\theta^{-1}-t} x \exp(-x(\theta^{-1}-t)) \right]_0^{+\infty} + \frac{\theta^{-2}}{\theta^{-1}-t} \int_0^{+\infty} \exp(-x(\theta^{-1}-t)) dx \\ = \frac{\theta^{-2}}{\theta^{-1}-t} \int_0^{+\infty} \exp(-x(\theta^{-1}-t)) dx \\ = \frac{\theta^{-2}}{\theta^{-1}-t} \left[\frac{-1}{\theta^{-1}-t} \exp(-x(\theta^{-1}-t)) \right]_0^{+\infty} = \frac{\theta^{-2}}{\theta^{-1}-t} \cdot \frac{1}{\theta^{-1}-t} \\ = \frac{\theta^{-2}}{(\theta^{-1}-t)^2} = \frac{\theta^{-2}}{\theta^{-2}(1-\theta t)^2} \\ = \frac{1}{(1-\theta t)^2}$$

$$\psi(t) = E[e^{tx}]; \quad \psi'(t) = E[x e^{tx}]; \quad \psi''(t) = E[x^2 e^{tx}]$$

$$\psi'(0) = E[x]; \quad \psi''(0) = E[x^2]$$

$$D_2, \quad \psi'(t) = 2\theta(1-\theta t)^{-3} \Rightarrow \psi'(0) = 2\theta \quad \psi''(t) = 6\theta^2(1-\theta t)^{-4} \\ = E[x] \Rightarrow \psi''(0) = 6\theta^2 = E[x^2]$$

